

BASIC RESULTS on PROBABILITY

Here we give results on probability that will be useful for Statistical Mechanics. Essentially this means we start from the theory for discrete events; this means that we will first need to recall basic facts about combinatorics (facts which are useful throughout physics). Later we go to results on correlations, on continuous variables, and so on. It is assumed here that everyone has already taken a course on probability - this document is largely intended to refresh your memory.

1. Permutations, Combinatorics, and Discrete Probabilities

We are interested in the theory of probability for a finite set of possible discrete outcomes. In physics this includes much of quantum mechanics, at least where a finite discrete set of states are involved; in statistical mechanics it involves the set of microstates for any quantum system. In more general areas of inquiry it can involve a huge range of activities and processes. Thus, the theory includes the calculation of the probability of any physical process which involves a finite set of outcomes. Well-known tutorial examples of this come from the calculation of probabilities for, eg., a possible outcome of a card game or indeed any game involving a finite set of possible states for the game.

One begins by considering the way in which N objects can be permuted amongst each other, ie., in how many different ways these N objects may be arranged (eg., on a line). There are two obvious cases here, viz.,

(i) The objects are all different or 'distinguishable' from each other. The number of different permutations is then simple to deduce. The object in the first position on the line may be chosen in N different ways, that in the second position in $N - 1$ ways, and so on. The number of possible arrangements is therefore $N(N - 1)(N - 2)\dots(1) = N!$

(ii) Alternatively, the objects are all indistinguishable. But then all the different permutations just given are identical - so there is only one way of arranging N indistinguishable objects on a line.

Multinomial Combinatorics: Suppose we now consider N objects of which we have n_1 of type 1, which are identical to each other, n_2 identical objects of type 2, etc., etc., up to n_m identical objects of type m (so that $\sum_j n_j = N$). As we have seen, if all the N objects were distinguishable, the number of their permutations

would be $N!$. However, the number of *distinguishable* permutations is only

$$C_{\{n_j\}}^N \equiv C_{n_1, n_2, \dots, n_m}^N = \frac{N!}{n_1! n_2! \dots n_m!} \quad (1)$$

since the j -th group of identical objects can be rearranged in $n_j!$ ways without changing anything, and we can do this for any of the m different sub-groups. The simplest example of a multinomial distribution is of course the binomial distribution, viz.,

$$C_n^N = \frac{N!}{n!(N-n)!} \quad (2)$$

The reason that these coefficients are given the name ‘multinomial’ coefficients is because they are just the coefficients in the ‘multinomial’ expansion of the expression $(z_1 + z_2 + \dots + z_m)^N$, given by

$$\left(\sum_{k=1}^m z_k \right)^N = \sum_{n_1} \sum_{n_2} \dots \sum_{n_m} \delta(N - \sum_k n_k) \left(\frac{N!}{n_1! n_2! \dots n_m!} \right) z_1^{n_1} z_2^{n_2} \dots z_m^{n_m} \quad (3)$$

Here the Kronecker delta function $\delta(N - \sum_k n_k)$ enforces the constraint that $\sum_k n_k = N$ (thus, eg., in our example above, the sum of the number of rocks in all the different piles must be equal to N).

Now there are various intuitive ways to think of these coefficients. Here are two:

(i) Imagine we have N identical balls which we distribute in m different cells or boxes. What then is the total number of different ways that this can be done, with n_1 balls in the 1st cell, n_2 in the 2nd, and so on. The answer is the multinomial distribution in (1).

(ii) Suppose I have n_1 balls of 1 colour (all then indistinguishable from each other), n_2 balls of another, and so on up to the m -th colour. How many different distinguishable ways can we then order the N balls? The answer is again the multinomial distribution in (1).

There are clearly many different variants on these. Thus, eg., suppose we have N distinguishable rocks which we divide into m piles, such that we have n_k objects in the k -th pile, with $k = i = 1, 2, \dots, m$. Suppose also that the ordering of rocks in each pile is irrelevant - they are just piles. Then it is clear that the multinomial coefficient gives the number of ways in which we can divide the different

distinguishable objects into these piles. To see this explicitly, let us imagine we start by picking the n_1 rocks for the first pile; we can clearly do this in $C_{n_1}^N$ ways. The n_2 rocks in the second pile can then be picked from the $(N - n_1)$ remaining ones in $C_{n_2}^{N-n_1}$ ways; and so on until we get to the 2nd last pile, ie., the $(m - 1)$ -th pile, and this can be done in $C_{N_{m-1}}^{N-n_1-\dots-n_{m-2}}$ ways. The remaining rocks, which make up the last pile, can only be found in one single way.

Multiplying all these together, we find that the total number of ways of dividing the original N rocks into m piles is given by

$$N = C_{n_1}^N C_{n_2}^{N-n_1} \dots C_{N_{m-1}}^{N-n_1-\dots-n_{m-2}} = \frac{N!}{n_1!n_2!\dots n_m!} \quad (4)$$

Here the last result is obtained by just multiplying out all the terms, and noting how the numerator in one such term is cancelled by a factor in the denominator of the previous term.

One can of course think of many other ways to count different orderings that involve these coefficients.

Discrete Probabilities: Now let us use these results for counting permutations to calculate probabilities for discrete outcomes (ie., where there is a finite number of possible outcomes for discrete events).

(a) In the first kind of problem we will consider, all of the different possible outcomes are assumed to be *a priori* equally probable. Usually it is physically obvious for real systems when this should be the case, in that there is nothing whatsoever that gives an y reason to suppose that one outcome is more likely than any other. A good example is the toss of a perfect coin, where by assumption the probability of getting heads (H) or tails (T) is equal (and therefore each has probability 1/2). In this case there are 4 possible outcomes if we toss two such coins, or if we toss the same coin twice, these being HH, HT, TH, and TT. Clearly each of these is equally likely, and so their probabilities are 1/4 each.

It is then clear that if we want to calculate the probability of one specific outcome when the total number of possible outcomes is Q , that probability must be $1/Q$. On the other hand, if we want to find the probability that we will get an outcome which itself involves a set of S different discrete outcomes (this set being a subset of all possible outcomes), then that probability will be S/Q .

This is where all the counting exercises we have done above come in handy. To see this, consider the probability that we will get n heads if we throw a perfect coin N times. The ordering is irrelevant, so that the total number of ways of throwing

n heads is just C_n^N , as we have already seen. However the total number of possible throws is clearly 2^N ; and so it follows that the probability $P_N(n)$ of getting n heads is just given by

$$P_N(n) = \frac{C_n^N}{2^N} = \frac{1}{2^N} \left(\frac{N!}{n!(N-n)!} \right) \quad (5)$$

When $N, n \gg 1$ we can easily find accurate expressions for this using Stirling's asymptotic formula.

(b) Now suppose that the probabilities for the different outcomes are *not* the same. At this point we must *assign* probabilities depending on what knowledge we have of the system involved. Suppose, eg., that for the coin discussed above, we know that the probability of getting heads is p_+ , so that the probability of getting tails is $p_- = (1 - p_+)$. We count things in the same way, but now we have to assign the correct probability to each outcome. It should be immediately obvious that the new result for $P_N(n)$ is now

$$P_N(n) = C_n^N p_+^n p_-^{N-n} \equiv \left(\frac{N!}{n!(N-n)!} \right) p_+^n (1 - p_+)^{N-n} \quad (6)$$

because the probability of getting any one of the combinations with n heads and $N - n$ tails is, by assumption, just $p_+^n p_-^{N-n}$. When $p_+ = 1/2$, this just reduces to the previous result.

It should now be obvious how to generalize this to cases where we have more than two different types of object involved in our outcomes, ie., where we deal with multinomial combinatorics. Thus, suppose we have N identical balls which we distribute in m different cells or boxes, but now the probability of going into the k -th box is p_k , where $k = 1, 2, \dots, m$ (and where of course $\sum_k p_k = 1$). As we saw before, the number of different ways of doing this is just the multinomial coefficient $C_{n_1, n_2, \dots, n_m}^N$; but now the weighting attached to any one of these ways is $\prod_k (p_k)^{n_k}$. It then follows that the probability $P_N(n_1, n_2, \dots, n_m)$ of getting an outcome in which there are n_k balls in the k -th box is just

$$\begin{aligned} P_N(n_1, n_2, \dots, n_m) &= C_{n_1, n_2, \dots, n_m}^N \prod_{k=1}^m p_k^{n_k} \\ &= \delta(N - \sum_k n_k) \left(\frac{N!}{n_1! n_2! \dots n_m!} \right) p_1^{n_1} p_2^{n_2} \dots p_m^{n_m} \quad (7) \end{aligned}$$

Again, we include the Kronecker delta constraint, as before. Note that if we now *sum* over all possible outcomes here (which means summing over all the different

values of the n_k within the constraint that $\sum_k n_k = N$), then we just get back the formula (3) above, with p_k substituted for z_k . Note that the left hand side of (3) then becomes unity, because $\sum_k p_k = 1$, and this is of course what we would expect - the sum of the probabilities of all different outcomes exhausts all possibilities, and so it must be unity.

Let's consider some examples of what we are talking about here, to give you an idea. I will, for simplicity, look here at cases where the probabilities of each outcome are all the same.

example 1: Suppose I draw 7 cards from a 52-card pack of cards. What is the probability that this hand of cards contains 3 Aces?

To do this we need to first ask how many possible outcomes there are for the 7 cards that are dealt; we then ask how many of these give 3 Aces. The probability is then the latter number divided by the former. The first question is simple - the total number of possible distinguishable arrangements is the binomial $C_7^{52} \equiv C_{45}^{52} = 52!/7! 45!$, because we can re-order the first 7 cards 7! times, and the last 45 cards 45! times.

To deal with the 2nd question we note first that it does not matter which Aces we get. We need to multiply the number of ways of getting 3 of the 4 Aces (without caring which ones), by the total number of outcomes for the other 4 cards that are dealt, with the constraint that these are NOT Aces. The first number is $C_3^4 = 4$. To find the second number, we note that there are 48 cards that are not Aces, and we are getting 4 of these. So this latter number is $C_4^{48} = 48!/44! 4!$.

The final result for the probability $P_{AAA}^{\{7\}}$ is then

$$P_{AAA}^{\{7\}} = \frac{C_3^4 C_4^{48}}{C_7^{52}} = 4 \times \frac{48!}{4! 44!} \times \frac{7! 45!}{52!} = 7.6.5.4 \frac{45}{52.51.50.49} \quad (8)$$

which if we work it out gives $P_{AAA}^{\{7\}} \sim 0.00582$, ie., roughly a probability of 1/172.

Example 2: In a game of poker, each of four players is dealt 5 cards from a pack of 52 cards. What is the probability that each player is dealt an ace?

A: This is a generalization of the last problem to a multinomial distribution. We must first ask how many possible outcomes there are for the 4 batches of 5 cards that are dealt; we then ask how many of these give 1 Ace in each hand. The probability is then the latter number divided by the former.

The answer to the first question is given by the multinomial distribution - we

have $C_{5.5.5.5.32}^{52} \equiv 52!/(5!)^4 32!$ ways of distributing the cards amongst 4 hands of 5 cards, and amongst the remaining 32 cards.

To deal with the second question we note first that the ordering of the Aces between the hands is irrelevant - there are $4!$ different ways of ordering the 4 Aces. There are then 48 cards left, that are not Aces - these can be dealt out to the 4 different hands in a total of $C_{4.4.4.4.32}^{48} \equiv 48!/(4!)^4 32!$ times.

The final result for the probability $P_{4A}^{\{4 \times 5\}}$ is then given by

$$P_{4A}^{\{4 \times 5\}} = \frac{4! \times C_{4.4.4.4.32}^{48}}{C_{5.5.5.5.32}^{52}} = \frac{5^4 \times 24}{52.51.50.49} \quad (9)$$

which, if we work it out, just gives $P_{4A}^{\{4 \times 5\}} \sim 2.31 \times 10^{-3}$, ie., a probability $\sim 1/433$.

Example 3: There is a British game called snooker, one of a large variety of different games of billiards. In this game, one has a white, a yellow, a green, a brown, a blue, a pink, a black and 15 red balls (for a total of 22 balls). So - how many different permutations can one make of these (ie., how many different distinguishable ways can one order them, assuming all the reds are **INDistinguishable** from each other)? And, if we pull balls out at random, what is the probability that the first 15 of these will be red?

A. We are clearly dealing with the multinomial distribution here, with the result that we have

$$\frac{22!}{(1!)^7 15!} \quad (10)$$

different distinguishable permutations. You can easily do the second part of this question yourself.